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Asymptotic theory for fractional regression models via Malliavin calculus

Solesne Bourguin ¹ Ciprian A. Tudor ^{2, *}

¹SAMM, Université de Paris 1 Panthéon-Sorbonne,
90, rue de Tolbiac, 75634, Paris, France.
solesne.bourguin@univ-paris1.fr

² Laboratoire Paul Painlevé, Université de Lille 1
F-59655 Villeneuve d'Ascq, France.
tudor@math.univ-lille1.fr

Abstract

We study the asymptotic behavior as $n \rightarrow \infty$ of the sequence

$$S_n = \sum_{i=0}^{n-1} K(n^\alpha B_i^{H_1}) \left(B_{i+1}^{H_2} - B_i^{H_2} \right)$$

where B^{H_1} and B^{H_2} are two independent fractional Brownian motions, K is a kernel function and the bandwidth parameter α satisfies certain hypotheses in terms of H_1 and H_2 . Its limiting distribution is a mixed normal law involving the local time of the fractional Brownian motion B^{H_1} . We use the techniques of the Malliavin calculus with respect to the fractional Brownian motion.

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Key words: limit theorems, fractional Brownian motion, multiple stochastic integrals, Malliavin calculus, regression model, weak convergence.

1 Introduction

The motivation of our work comes from the econometric theory. Consider a regression model of the form

$$y_i = f(x_i) + u_i, \quad i \geq 0$$

*Associate member of the team Samm, Université de Panthéon-Sorbonne Paris 1

where $(u_i)_{i \geq 0}$ is the "error" and $(x_i)_{i \geq 0}$ is the regressor. The purpose is to estimate the function f based on the observation of the random variables y_i , $i \geq 0$. The conventional kernel estimate of $f(x)$ is

$$\hat{f}(x) = \frac{\sum_{i=0}^n K_h(x_i - x) y_i}{\sum_{i=0}^n K_h(x_i - x)}$$

where K is a nonnegative real kernel function satisfying $\int_{\mathbb{R}} K^2(y) dy = 1$ and $\int_{\mathbb{R}} y K(y) dy = 0$ and $K_h(s) = \frac{1}{h} K(\frac{s}{h})$. The bandwidth parameter $h \equiv h_n$ satisfies $h_n \rightarrow 0$ as $n \rightarrow \infty$. The asymptotic behavior of the estimator \hat{f} is usually related to the behavior of the sequence

$$V_n = \sum_{i=1}^n K_h(x_i - x) u_i.$$

The limit in distribution as $n \rightarrow \infty$ of the sequence S_n has been widely studied in the literature in various situations. We refer, among others, to [8] and [9] for the case where x_t is a recurrent Markov chain, to [15] for the case where x_t is a partial sum of a general linear process, and [16] for a more general situation. See also [13] or [14]. An important assumption in the main part of the above references is the fact that u_i is a martingale difference sequence. In our work we will consider the following situation: we assume that the regressor $x_i = B_i^{H_1}$ is a fractional Brownian motion (fBm) with Hurst parameter $H_1 \in (0, 1)$ and the error is $u_i = B_{i+1}^{H_2} - B_i^{H_2}$ where B^{H_2} is a fBm with $H_2 \in (0, 1)$ and it is independent from B^{H_1} . In this case, our error process has no semimartingale property. We will also set $h_n = n^{-\alpha}$ with $\alpha > 0$. A supplementary assumption on α will be imposed later in terms of the Hurst parameters H_1 and H_2 . The sequence V_n can be now written as

$$S_n(x) = \sum_{i=0}^n K(n^\alpha(B_i^{H_1} - x)) (B_{i+1}^{H_2} - B_i^{H_2}). \quad (1)$$

Our purpose is to give an approach based on stochastic calculus for this asymptotic theory. Recently, the stochastic integration with respect to the fractional Brownian motion has been widely studied. Various types of stochastic integrals, based on Malliavin calculus, Wick products or rough path theory have been introduced and change of variables formulas have been derived. We will use all these different techniques in our work. The general idea is as follows. Suppose that $x = 0$. We will first observe that the asymptotic behavior of the sequence S_n will be given by the sum

$$a_n = \sum_{i,j=0}^n K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \mathbf{E} \left((B_{i+1}^{H_2} - B_i^{H_2}) (B_{j+1}^{H_2} - B_j^{H_2}) \right). \quad (2)$$

This is easy to understand since the conditional distribution of S_n given B^{H_1} is given by

$$(a_n)^{\frac{1}{2}} Z$$

where Z is a standard normal random variable. The double sum a_n can be decomposed into two parts: a "diagonal" part given by $\sum_{i=1}^n K^2(n^\alpha B_i^{H_1})$ and a "non-diagonal" part given

by the terms with $i \neq j$. We will restrict ourselves to the situation where the diagonal part is dominant (in a sense that will be defined later) with respect to the non-diagonal part. This will imply a certain assumption on the bandwidth parameter α in terms of H_1 and H_2 . We will therefore need to study the asymptotic behavior of

$$\langle S \rangle_n := \sum_{i=1}^n K^2(n^\alpha B_i^{H_1}). \quad (3)$$

(In the case $H_2 = \frac{1}{2}$ this is actually the bracket of S_n which is a martingale; this motivates our choice of notation.) We will assume that the kernel K is the standard Gaussian kernel

$$K(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}.$$

This choice is motivated by the fact that $K^2(n^\alpha B_i^{H_1})$ can be decomposed into an orthogonal sum of multiple Wiener-Itô integrals (see [11], [3], [4]) and the Malliavin calculus can be used to treat the convergence of (3). Its limit in distribution will be after normalization the local time of the fractional Brownian motion denoted $cL^{H_1}(1,0)$, where c is positive constant. Consequently, we will find that the (renormalized) sequence S_n converges in law to a mixed normal random variable $cW_{L^{H_1}(1,0)}$ where W is a Brownian motion independent from B^{H_1} and c is a positive constant. The result is in concordance with the papers [15], [16].

But we also prove a stronger result: we show that the vector $(S_n, (G_t)_{t \geq 0})$ converges in the sense of finite dimensional distributions to the vector $(cW_{L^{H_1}(1,0)}, (G_t)_{t \geq 0})$, where c is a positive constant, for any stochastic process $(G_t)_{t \geq 0}$ independent from B^{H_1} and adapted to the filtration generated by B^{H_2} which satisfies some regularity properties in terms of the Malliavin calculus. We will say that S_n converges stably to its limit. To prove this stable convergence we will express S_n as a stochastic integral with respect to B^{H_2} and we will use the techniques of the Malliavin calculus. We will limit ourselves in this last section to the case $H_2 > \frac{1}{2}$.

We also mention that, although the error process B^{H_2} does not appear in the limit of (1), it governs the behavior of this sequence. Indeed, the parameter H_2 is involved in the renormalization of (1) and the stochastic calculus with respect to B^{H_2} is crucial in the proof of our main results.

We have organized our paper as follows: Section 2 contains the notations, definitions and results from the stochastic calculus that will be needed throughout our paper. In Section 3 we will find the renormalization order of the sequence (1), while Section 4 contains the result on the convergence of the “bracket” (3). In Section 5 we will prove the limit theorem in distribution for $S_n(0)$ and in Section 6 we will discuss the stable convergence of this sequence.

2 Preliminaries

Here we describe the elements from stochastic analysis that we will need in the paper. Consider \mathcal{H} a real separable Hilbert space and $(B(\varphi), \varphi \in \mathcal{H})$ an isonormal Gaussian process on a probability space (Ω, \mathcal{A}, P) , that is a centered Gaussian family of random variables such that $\mathbf{E}(B(\varphi)B(\psi)) = \langle \varphi, \psi \rangle_{\mathcal{H}}$. Denote by I_n the multiple stochastic integral with respect to B (see [10]). This I_n is actually an isometry between the Hilbert space $\mathcal{H}^{\odot n}$ (symmetric tensor product) equipped with the scaled norm $\frac{1}{\sqrt{n!}} \|\cdot\|_{\mathcal{H}^{\otimes n}}$ and the Wiener chaos of order n which is defined as the closed linear span of the random variables $H_n(B(\varphi))$ where $\varphi \in \mathcal{H}$, $\|\varphi\|_{\mathcal{H}} = 1$ and H_n is the Hermite polynomial of degree $n \geq 1$

$$H_n(x) = \frac{(-1)^n}{n!} \exp\left(\frac{x^2}{2}\right) \frac{d^n}{dx^n} \left(\exp\left(-\frac{x^2}{2}\right) \right), \quad x \in \mathbb{R}.$$

The isometry of multiple integrals can be written as: for m, n positive integers,

$$\begin{aligned} \mathbf{E}(I_n(f)I_m(g)) &= n! \langle f, g \rangle_{\mathcal{H}^{\otimes n}} \quad \text{if } m = n, \\ \mathbf{E}(I_n(f)I_m(g)) &= 0 \quad \text{if } m \neq n. \end{aligned} \tag{4}$$

It also holds that

$$I_n(f) = I_n(\tilde{f})$$

where \tilde{f} denotes the symmetrization of f defined by $\tilde{f}(x_1, \dots, x_n) = \frac{1}{n!} \sum_{\sigma \in \mathcal{S}_n} f(x_{\sigma(1)}, \dots, x_{\sigma(n)})$.

We recall that any square integrable random variable which is measurable with respect to the σ -algebra generated by B can be expanded into an orthogonal sum of multiple stochastic integrals

$$F = \sum_{n \geq 0} I_n(f_n) \tag{5}$$

where $f_n \in \mathcal{H}^{\odot n}$ are (uniquely determined) symmetric functions and $I_0(f_0) = \mathbf{E}[F]$.

Let L be the Ornstein-Uhlenbeck operator

$$LF = - \sum_{n \geq 0} n I_n(f_n)$$

if F is given by (5).

For $p > 1$ and $\alpha \in \mathbb{R}$ we introduce the Sobolev-Watanabe space $\mathbb{D}^{\alpha, p}$ as the closure of the set of polynomial random variables with respect to the norm

$$\|F\|_{\alpha, p} = \|(I - L)^{\frac{\alpha}{2}}\|_{L^p(\Omega)}$$

where I represents the identity. We denote by D the Malliavin derivative operator that acts on smooth functions of the form $F = g(B(\varphi_1), \dots, B(\varphi_n))$ (g is a smooth function with compact support and $\varphi_i \in \mathcal{H}$)

$$DF = \sum_{i=1}^n \frac{\partial g}{\partial x_i}(B(\varphi_1), \dots, B(\varphi_n)) \varphi_i.$$

The operator D is continuous from $\mathbb{D}^{\alpha,p}$ into $\mathbb{D}^{\alpha-1,p}(\mathcal{H})$. The adjoint of D is denoted by δ and is called the divergence (or Skorohod) integral. It is a continuous operator from $\mathbb{D}^{\alpha,p}(\mathcal{H})$ into $\mathbb{D}^{\alpha-1,p}$. We have the following duality relationship between D and δ

$$\mathbf{E}(F\delta(u)) = \mathbf{E}\langle DF, u \rangle_{\mathcal{H}} \text{ for every } F \text{ smooth.} \quad (6)$$

For adapted integrands, the divergence integral coincides with the classical Itô integral. We will use the notation

$$\delta(u) = \int_0^T u_s dB_s.$$

Let u be a stochastic process having the chaotic decomposition $u_s = \sum_{n \geq 0} I_n(f_n(\cdot, s))$ where $f_n(\cdot, s) \in \mathcal{H}^{\otimes n}$ for every s . One can prove that $u \in \text{Dom } \delta$ if and only if $\tilde{f}_n \in \mathcal{H}^{\otimes(n+1)}$ for every $n \geq 0$, and $\sum_{n=0}^{\infty} I_{n+1}(\tilde{f}_n)$ converges in $L^2(\Omega)$. In this case,

$$\delta(u) = \sum_{n=0}^{\infty} I_{n+1}(\tilde{f}_n) \quad \text{and} \quad \mathbf{E}|\delta(u)|^2 = \sum_{n=0}^{\infty} (n+1)! \|\tilde{f}_n\|_{\mathcal{H}^{\otimes(n+1)}}^2.$$

In our work we will mainly consider divergence integrals with respect to a fractional Brownian motion. The fractional Brownian motion $(B_t^H)_{t \in [0, T]}$ with Hurst parameter $H \in (0, 1)$ is a centered Gaussian process starting from zero with covariance function

$$R^H(t, s) := \frac{1}{2} (t^{2H} + s^{2H} - |t - s|^{2H}), \quad s, t \in [0, T].$$

In this case the space $\mathcal{H}_{\mathcal{H}}$ is the canonical Hilbert space of the fractional Brownian motion which is defined as the closure of the linear space generated by the indicator functions $\{1_{[0, t]}, t \in [0, T]\}$ with respect to the scalar product

$$\langle 1_{[0, t]}, 1_{[0, s]} \rangle_{\mathcal{H}_H} = R^H(t, s), \quad s, t \in [0, T].$$

3 Renormalization of the sequence S_n

As we mentioned in the introduction, we will assume throughout the paper that $x = 0$ in (1), then

$$S_n := S_n(0) = \sum_{i=0}^{n-1} K(n^\alpha B_i^{H_1})(B_{i+1}^{H_2} - B_i^{H_2}). \quad (7)$$

We compute in this part the L^2 -norm of S_n in order to renormalize it. We have

$$\begin{aligned}
\mathbf{E}(S_n^2) &= \mathbf{E} \left(\sum_{i,j=0}^{n-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) (B_{i+1}^{H_2} - B_i^{H_2}) (B_{j+1}^{H_2} - B_j^{H_2}) \right) \\
&= \mathbf{E} \left(\sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) (B_{i+1}^{H_2} - B_i^{H_2})^2 \right) \\
&\quad + \mathbf{E} \left(\sum_{i \neq j}^{n-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) (B_{i+1}^{H_2} - B_i^{H_2}) (B_{j+1}^{H_2} - B_j^{H_2}) \right) \\
&= T' + T''.
\end{aligned}$$

The summand T' will be called the “diagonal” term while the summand T'' will be called “the non-diagonal” term. We will analyze each of them separately. Concerning T' we have

Lemma 1 *As $n \rightarrow +\infty$,*

$$n^{\alpha+H_1-1} T' \xrightarrow{n \rightarrow +\infty} C_1 = \frac{1}{2\pi\sqrt{2}(1-H_1)}. \quad (8)$$

Proof: Through the independence of $(B_t^{H_1})_{t \geq 0}$ and $(B_t^{H_2})_{t \geq 0}$,

$$T' = \sum_{i=0}^{n-1} \mathbf{E} \left(K^2(n^\alpha B_i^{H_1}) \right) \mathbf{E} \left((B_{i+1}^{H_2} - B_i^{H_2})^2 \right).$$

Since $\mathbf{E} \left((B_{i+1}^{H_2} - B_i^{H_2})^2 \right) = 1$,

$$T' = \sum_{i=0}^{n-1} \mathbf{E} \left(K^2(n^\alpha B_i^{H_1}) \right) = \sum_{i=0}^{n-1} \mathbf{E} \left(\frac{1}{2\pi} e^{-n^{2\alpha} i^{2H_1} Z^2} \right)$$

where Z is a standard normal random variable. Recall that, if Z is a standard normal random variable, and if $1 + 2c > 0$

$$\mathbf{E} \left(e^{-cZ^2} \right) = \frac{1}{\sqrt{1+2c}} \quad (9)$$

consequently,

$$T' = \sum_{i=0}^{n-1} \frac{1}{2\pi\sqrt{1+2n^{2\alpha} i^{2H_1}}}.$$

As $n \rightarrow +\infty$, T' behaves as such

$$\begin{aligned} \sum_{i=0}^{n-1} \frac{1}{2\pi\sqrt{1+2n^{2\alpha}i^{2H_1}}} &\sim \frac{n^{-\alpha}}{2\pi\sqrt{2}} \sum_{i=0}^{n-1} i^{-H_1} \sim \frac{n^{-\alpha-H_1+1}}{2\pi\sqrt{2}} \frac{1}{n} \sum_{i=0}^{n-1} \left(\frac{i}{n}\right)^{-H_1} \\ &\sim \frac{n^{-\alpha-H_1+1}}{2\pi\sqrt{2}} \int_0^1 x^{-H_1} dx = \frac{n^{-\alpha-H_1+1}}{2\pi\sqrt{2}(1-H_1)}. \end{aligned}$$

The sign “ \sim ” means that the left-hand side and the right-hand side have the same limit as $n \rightarrow +\infty$. We will use this notation throughout the paper. \blacksquare

We will now compute the term T'' . To do so, we will need the following Lemma (lemma 3.1 p. 122 in [17]).

Lemma 2 *For every $s, r \in [0, T]$, $s \geq r$ and $0 < H < 1$ we have*

$$s^{2H} r^{2H} - \mu^2 \geq \tau(s-r)^{2H} r^{2H} \quad (10)$$

where $\mu = \mathbf{E}(B_s^H B_r^H)$ and $\tau > 0$ is a constant.

Concerning the non-diagonal term of $\mathbf{E}(S_n^2)$ the following holds

Lemma 3 *Suppose that*

$$\alpha - 4H_2 + H_1 + 2 > 0. \quad (11)$$

Then, as $n \rightarrow +\infty$,

$$n^{\alpha+H_1-1} T'' \xrightarrow{n \rightarrow +\infty} 0. \quad (12)$$

Proof: Using again the independence of $(B_t^{H_1})_{t \geq 0}$ and $(B_t^{H_2})_{t \geq 0}$

$$\begin{aligned} T'' &= \sum_{i \neq j}^{n-1} \mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) \mathbf{E} \left((B_{i+1}^{H_2} - B_i^{H_2})(B_{j+1}^{H_2} - B_j^{H_2}) \right) \\ &= \frac{1}{2} \sum_{i \neq j}^{n-1} \mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) f_{H_2}(i, j) \end{aligned}$$

where

$$f_{H_2}(i, j) = \frac{1}{2} \left[|i-j+1|^{2H_2} + |i-j-1|^{2H_2} - 2|i-j|^{2H_2} \right]. \quad (13)$$

We need to evaluate the expectation $\mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right)$. Let $\Gamma = \begin{pmatrix} i^{2H_1} & R(i, j) \\ R(i, j) & j^{2H_1} \end{pmatrix}$ be the covariance matrix of $(B_i^{H_1}, B_j^{H_1})$. We have $|\Gamma| = (ij)^{2H_1} - R^2(i, j)$ and $\Gamma^{-1} = \frac{1}{|\Gamma|} \begin{pmatrix} j^{2H_1} & -R(i, j) \\ -R(i, j) & i^{2H_1} \end{pmatrix}$. The density of $(B_i^{H_1}, B_j^{H_1})$ is then

$$f(x, y) = \frac{1}{2\pi\sqrt{|\Gamma|}} e^{-\frac{1}{2|\Gamma|}(j^{2H_1}x^2 - 2R(i, j)xy + i^{2H_1}y^2)}. \quad (14)$$

We obtain

$$\begin{aligned}
\mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) &= \frac{1}{(2\pi)^2 \sqrt{|\Gamma|}} \int_{\mathbb{R}^2} e^{-\frac{n^{2\alpha} x^2}{2}} e^{-\frac{n^{2\alpha} y^2}{2}} e^{-\frac{1}{2|\Gamma|} (j^{2H_1} x^2 - 2R(i,j)xy + i^{2H_1} y^2)} dx dy \\
&= \frac{1}{(2\pi)^2 \sqrt{|\Gamma|}} \int_{\mathbb{R}} e^{-\frac{n^{2\alpha} y^2}{2}} e^{-\frac{i^{2H_1} y^2}{2|\Gamma|}} \int_{\mathbb{R}} e^{-\frac{n^{2\alpha} x^2}{2}} e^{-\frac{1}{2|\Gamma|} (j^{2H_1} x^2 - 2R(i,j)xy)} dx dy \\
&= \frac{1}{(2\pi)^2 \sqrt{|\Gamma|}} \int_{\mathbb{R}} e^{-\frac{y^2}{2} \left[n^{2\alpha} + \frac{i^{2H_1}}{|\Gamma|} \right]} \int_{\mathbb{R}} e^{-\frac{1}{2} \left[x^2 \left(n^{2\alpha} + \frac{j^{2H_1}}{|\Gamma|} \right) - \frac{2R(i,j)}{|\Gamma|} xy \right]} dx dy \\
&= \frac{1}{(2\pi)^2 \sqrt{|\Gamma|}} \int_{\mathbb{R}} e^{-\frac{y^2}{2} \left[n^{2\alpha} + \frac{i^{2H_1}}{|\Gamma|} \right]} \int_{\mathbb{R}} e^{-\frac{\left(n^{2\alpha} + \frac{j^{2H_1}}{|\Gamma|} \right)}{2} \left[x^2 - \frac{2R(i,j)}{n^{2\alpha} |\Gamma| + j^{2H_1}} xy \right]} dx dy \\
&= \frac{1}{(2\pi)^2 \sqrt{|\Gamma|}} \int_{\mathbb{R}} e^{-\frac{y^2}{2} \left[n^{2\alpha} + \frac{i^{2H_1}}{|\Gamma|} \right]} \int_{\mathbb{R}} e^{-\frac{\left(n^{2\alpha} + \frac{j^{2H_1}}{|\Gamma|} \right)}{2} \left[\left(x - \frac{R(i,j)}{n^{2\alpha} |\Gamma| + j^{2H_1}} y \right)^2 - \frac{R^2(i,j)}{(n^{2\alpha} |\Gamma| + j^{2H_1})^2} y^2 \right]} dx dy \\
&= \frac{1}{(2\pi)^2 \sqrt{|\Gamma|}} \int_{\mathbb{R}} e^{-\frac{y^2}{2} \left[n^{2\alpha} + \frac{i^{2H_1}}{|\Gamma|} \right]} e^{-\frac{\left(n^{2\alpha} + \frac{j^{2H_1}}{|\Gamma|} \right)}{2} \frac{R^2(i,j) y^2}{\left(n^{2\alpha} + \frac{j^{2H_1}}{|\Gamma|} \right)^2 |\Gamma|^2}} \int_{\mathbb{R}} e^{-\frac{\left(n^{2\alpha} + \frac{j^{2H_1}}{|\Gamma|} \right)}{2} \left(x - \frac{R(i,j)}{n^{2\alpha} |\Gamma| + j^{2H_1}} y \right)^2} dx dy \\
&= \frac{1}{(2\pi)^{\frac{3}{2}} \sqrt{|\Gamma|}} \frac{\sqrt{|\Gamma|}}{\sqrt{n^{2\alpha} |\Gamma| + j^{2H_1}}} \int_{\mathbb{R}} e^{-\frac{1}{2} y^2 \left[\frac{(n^{2\alpha} |\Gamma| + i^{2H_1})(n^{2\alpha} |\Gamma| + j^{2H_1}) - R^2(i,j)}{|\Gamma| (n^{2\alpha} |\Gamma| + j^{2H_1})} \right]} dy.
\end{aligned}$$

Thus

$$\begin{aligned}
\mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) &= \frac{1}{2\pi \sqrt{n^{2\alpha} |\Gamma| + j^{2H_1}}} \frac{\sqrt{|\Gamma|} \sqrt{(n^{2\alpha} |\Gamma| + j^{2H_1})}}{\sqrt{(n^{2\alpha} |\Gamma| + i^{2H_1})(n^{2\alpha} |\Gamma| + j^{2H_1}) - R^2(i,j)}} \\
&= \frac{\sqrt{|\Gamma|}}{2\pi \sqrt{(n^{2\alpha} |\Gamma| + i^{2H_1})(n^{2\alpha} |\Gamma| + j^{2H_1}) - R^2(i,j)}} \\
&= \frac{1}{2\pi \sqrt{n^{4\alpha} |\Gamma| + n^{2\alpha} j^{2H_1} + n^{2\alpha} i^{2H_1} + 1}}.
\end{aligned}$$

Suppose that $i > j$. We use Lemma 2 to bound $|\Gamma| = i^{2H_1} j^{2H_1} - R^2(i, j)$ from below. Therefore

$$\mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) \leq \frac{1}{2\pi \sqrt{n^{4\alpha} \tau (i-j)^{2H_1} j^{2H_1} + n^{2\alpha} (i^{2H_1} + j^{2H_1})}}.$$

Since $a^2 + b^2 \geq 2ab$ with $a^2 = n^{4\alpha} \tau (i-j)^{2H_1} j^{2H_1}$ and $b^2 = n^{2\alpha} (i^{2H_1} + j^{2H_1})$

$$\mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) \leq \frac{1}{2\pi \sqrt{2\sqrt{\tau} n^{2\alpha} (i-j)^{H_1} j^{H_1} \sqrt{n^{2\alpha} (i^{2H_1} + j^{2H_1})}}}$$

and using the same inequality as above for $a^2 = i^{2H_1}$ and $b^2 = j^{2H_1}$

$$\mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) \leq \frac{n^{-\frac{3\alpha}{2}}}{2\pi\sqrt{2}\tau^{\frac{1}{4}}(i-j)^{\frac{H_1}{2}} j^{\frac{3H_1}{4}} i^{\frac{H_1}{4}}}. \quad (15)$$

Since $f_{H_2}(i, j)$ behaves as $H_2(2H_2 - 1)|i - j|^{2H_2-2}$ when $i - j \rightarrow \infty$, we can assert that

$$T'' \sim \frac{H_2(2H_2 - 1)}{2} \sum_{i \neq j}^{n-1} \mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) |i - j|^{2H_2-2}.$$

Using (15), we can write

$$\sum_{i \neq j}^{n-1} \mathbf{E} \left(K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \right) |i - j|^{2H_2-2} \lesssim \sum_{i > j}^{n-1} \frac{n^{-\frac{3\alpha}{2}}}{2\pi\sqrt{2}\tau^{\frac{1}{4}}(i-j)^{\frac{H_1}{2}} j^{\frac{3H_1}{4}} i^{\frac{H_1}{4}}} |i - j|^{2H_2-2}$$

and consequently

$$\begin{aligned} T'' &\lesssim \frac{H_2(2H_2 - 1)}{4\pi\sqrt{2}\tau^{\frac{1}{4}}} n^{-\frac{3\alpha}{2}} n^{2H_2 - \frac{H_1}{2} - 2} n^{-\frac{3H_1}{4}} n^{-\frac{H_1}{4}} n^2 \underbrace{\frac{1}{n^2} \sum_{i > j}^{n-1} \frac{\left(\frac{i-j}{n}\right)^{2H_2 - \frac{H_1}{2} - 2}}{\left(\frac{j}{n}\right)^{\frac{3H_1}{4}}} \left(\frac{j}{n}\right)^{\frac{H_1}{4}}}_{\xrightarrow{n \rightarrow +\infty} C(H_1, H_2) > 0} \\ &\lesssim \frac{H_2(2H_2 - 1)C(H_1, H_2)}{4\pi\sqrt{2}\tau^{\frac{1}{4}}} n^{-\frac{3\alpha}{2} + 2H_2 - \frac{3H_1}{2}}. \end{aligned} \quad (16)$$

It follows that under condition (11) $n^{\alpha+H_1-1}T''$ converges to zero as $n \rightarrow \infty$. ■

As a consequence of Lemmas 1 and 3 we obtain the following L^2 - norm estimate for S_n .

Proposition 1 *Suppose that condition (11) holds. Then, as $n \rightarrow \infty$*

$$n^{\alpha+H_1-1} \mathbf{E} (S_n^2) \rightarrow C_1 = \frac{1}{2\pi\sqrt{2}(1-H_1)}.$$

The condition (11) will be discussed more thoroughly later (Remark 1, Section 5).

4 The limit in distribution of $\langle S \rangle_n$

Proposition 1 implies that the diagonal part of S_n^2 is dominant in relation to the non-diagonal part, in the sense that this diagonal part is responsible for the renormalization order of S_n^2 which is $n^{\alpha+H_1-1}$. As a consequence we need to study the limit distribution

of $n^{\alpha+H_1-1}\langle S \rangle_n = n^{\alpha+H_1-1} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})$. Using the self-similarity property of the fractional Brownian motion we have

$$n^{\alpha+H_1-1} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) = n^{\alpha+H_1-1} \sum_{i=0}^{n-1} K^2(n^{\alpha+H_1} B_{\frac{i}{n}}^{H_1}).$$

The limit of the above sequence is linked to the local time of the fractional Brownian motion B^{H_1} . For any $t \geq 0$ and $x \in \mathbb{R}$ we define $L^{H_1}(t, x)$ as the density of the occupation measure (see [1], [5])

$$\mu_t(A) = \int_0^t 1_A(B_s^{H_1}) ds, \quad A \in \mathcal{B}(\mathbb{R}).$$

The local time $L^{H_1}(t, x)$ satisfies the occupation time formula

$$\int_0^t f(B_s^{H_1}) ds = \int_{\mathbb{R}} L^{H_1}(t, x) f(x) dx \quad (17)$$

for any measurable function f . The local time is Hölder continuous with respect to t and with respect to x (for the sake of completeness $L^{H_1}(t, x)$ has Hölder continuous paths of order $\delta < 1 - H$ in time and of order $\gamma < \frac{1-H}{2H}$ in the space variable (see Table 2 in [5])). Moreover, it admits a bicontinuous version with respect to (t, x) .

Below, we give an important convergence result that will be necessary in proving the main result of this section.

Proposition 2 *The following convergence in distribution result holds*

$$n^{\alpha+H_1} \left(\frac{1}{n} \sum_{i=0}^{n-1} K^2(n^{\alpha+H_1} B_{\frac{i}{n}}^{H_1}) - \int_0^1 K^2(n^{\alpha+H_1} B_s^{H_1}) ds \right) \xrightarrow{n \rightarrow +\infty} 0. \quad (18)$$

Proof: Fix $\varepsilon > 0$. Let $p_\varepsilon(x)$ be the Gaussian kernel with variance $\varepsilon > 0$ defined by $p_\varepsilon(x) = \frac{1}{\sqrt{2\pi\varepsilon}} e^{-\frac{x^2}{2\varepsilon}}$. Note that for every $s \geq 0$

$$\sqrt{\pi} n^{\alpha+H_1} K^2(n^{\alpha+H_1} B_s^{H_1}) = \frac{1}{2} p_{\frac{1}{2n^{2(\alpha+H_1)}}}(B_s^{H_1}). \quad (19)$$

Using (19), we can write the left-hand side of (18) as

$$\begin{aligned}
& \sqrt{\pi} n^{\alpha+H_1} \left(\int_0^1 K^2(n^{\alpha+H_1} B_s^{H_1}) ds - \frac{1}{n} \sum_{i=0}^{n-1} K^2(n^{\alpha+H_1} B_{\frac{i}{n}}^{H_1}) \right) \\
&= \frac{1}{2} \sum_{i=0}^{n-1} \int_{\frac{i}{n}}^{\frac{i+1}{n}} \left(p_{\frac{1}{2}n^{-2(\alpha+H_1)}}(B_s^{H_1}) - p_{\frac{1}{2}n^{-2(\alpha+H_1)}}(B_{\frac{i}{n}}^{H_1}) \right) ds \\
&= \frac{1}{2} \sum_{i=0}^{n-1} \int_{\frac{i}{n}}^{\frac{i+1}{n}} \left(p_{\frac{1}{2}n^{-2(\alpha+H_1)}}(B_s^{H_1}) - p_\varepsilon(B_s^{H_1}) \right) ds \\
&\quad + \frac{1}{2} \sum_{i=0}^{n-1} \int_{\frac{i}{n}}^{\frac{i+1}{n}} \left(p_\varepsilon(B_s^{H_1}) - p_\varepsilon(B_{\frac{i}{n}}^{H_1}) \right) ds \\
&\quad + \frac{1}{2} \sum_{i=0}^{n-1} \int_{\frac{i}{n}}^{\frac{i+1}{n}} \left(p_\varepsilon(B_{\frac{i}{n}}^{H_1}) - p_{\frac{1}{2}n^{-2(\alpha+H_1)}}(B_{\frac{i}{n}}^{H_1}) \right) ds := \frac{1}{2}(a_n^{(1)} + a_n^{(2)} + a_n^{(3)}).
\end{aligned}$$

We will now estimate the three terms above and we will show that each of them converges to zero (in some sense). Let us first handle the term $a_n^{(1)}$. We have

$$a_n^{(1)} = \int_0^1 p_{\frac{1}{2}n^{-2(\alpha+H_1)}}(B_s^{H_1}) ds - \int_0^1 p_\varepsilon(B_s^{H_1}) ds.$$

It follows from [11] or [4] that

$$\int_0^1 p_\varepsilon(B_s^{H_1}) ds \xrightarrow{\varepsilon \rightarrow 0} \int_0^1 \delta_0(B_s^{H_1}) ds = L^{H_1}(1, 0) \quad (20)$$

in $L^2(\Omega)$ and almost surely, where $L^{H_1}(1, 0)$ is the local time of the fractional Brownian motion. Therefore $a_n^{(1)}$ clearly converges to zero as $\varepsilon \rightarrow 0$ and $n \rightarrow \infty$. The term $a_n^{(2)}$ can be expressed as

$$a_n^{(2)} = - \left(\frac{1}{n} \sum_{i=0}^{n-1} p_\varepsilon(B_{\frac{i}{n}}^{H_1}) - \int_0^1 p_\varepsilon(B_s^{H_1}) ds \right) \quad (21)$$

and for every $\varepsilon > 0$ it converges almost surely to zero as $n \rightarrow \infty$ using the Riemann sum convergence. Let us now handle the term $a_n^{(3)}$ given by

$$a_n^{(3)} = \frac{1}{n} \sum_{i=0}^{n-1} \left(p_\varepsilon(B_{\frac{i}{n}}^{H_1}) - p_{\frac{1}{2}n^{-2(\alpha+H_1)}}(B_{\frac{i}{n}}^{H_1}) \right). \quad (22)$$

We will treat this term by using the chaos decomposition of the Gaussian kernel applied to random variables in the first Wiener chaos. Recall that (see [3], [6], [7], [12]) for every $\varphi \in \mathcal{H}_{H_1}$ (\mathcal{H}_{H_1} is the canonical Hilbert space associated with the Gaussian process B^{H_1}),

$$p_\varepsilon(B^{H_1}(\varphi)) = \sum_{m \geq 0} C_m I_{2m}(\varphi^{\otimes 2m}) \frac{1}{(\|\varphi\|_{\mathcal{H}_1}^2 + \varepsilon)^{m + \frac{1}{2}}} \quad (23)$$

where $C_m = \frac{(-1)^m}{\sqrt{2\pi}2^m m!}$.

Using this chaos decomposition, we can write $p_\varepsilon(B_{\frac{i}{n}}^{H_1}) - p_{\frac{1}{2}n^{-2(\alpha+H_1)}}(B_{\frac{i}{n}}^{H_1})$ as

$$\begin{aligned} p_\varepsilon(B_{\frac{i}{n}}^{H_1}) - p_{\frac{1}{2}n^{-2(\alpha+H_1)}}(B_{\frac{i}{n}}^{H_1}) &= \sum_{m \geq 0} C_m I_{2m} \left(1_{[0, \frac{i}{n}]}^{\otimes 2m} \right) \left(\frac{1}{\left(\left(\frac{i}{n} \right)^{2H_1} + \varepsilon \right)^{m+\frac{1}{2}}} - \frac{1}{\left(\left(\frac{i}{n} \right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)} \right)^{m+\frac{1}{2}}} \right) \\ &= \sum_{m \geq 0} C_m I_{2m} \left(1_{[0, \frac{i}{n}]}^{\otimes 2m} \right) \left(\frac{i}{n} \right)^{-2H_1(m+\frac{1}{2})} d_{i,\varepsilon,n,m} \end{aligned}$$

where

$$d_{i,\varepsilon,n,m} = \left(\left(\frac{\left(\frac{i}{n} \right)^{2H_1}}{\left(\left(\frac{i}{n} \right)^{2H_1} + \varepsilon \right)} \right)^{m+\frac{1}{2}} - \left(\frac{\left(\frac{i}{n} \right)^{2H_1}}{\left(\left(\frac{i}{n} \right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)} \right)} \right)^{m+\frac{1}{2}} \right).$$

We will show that $a_n^{(3)}$ converges to zero in $L^2(\Omega)$ as $n \rightarrow \infty$ and $\varepsilon \rightarrow 0$. From (22) one can easily see that the diagonal part of $a_n^{(3)}$ converges to zero. We can also see, from the expression of $a_n^{(3)}$, that the summands with $j = 0$ vanish. Then, by using the orthogonality of multiple stochastic integrals([10]), we obtain

$$\mathbf{E}(a_n^{(3)})^2 \sim \frac{1}{n^2} \sum_{m \geq 0} C_m^2 (2m)! \sum_{i,j \geq 1; i \neq j}^{n-1} \langle 1_{[0, \frac{i}{n}]}, 1_{[0, \frac{j}{n}]} \rangle_{\mathcal{H}_1}^{2m} \left(\frac{i}{n} \right)^{-2H_1(m+\frac{1}{2})} \left(\frac{j}{n} \right)^{-2H_1(m+\frac{1}{2})} d_{i,\varepsilon,n,m} d_{j,\varepsilon,n,m}.$$

We can also write

$$\begin{aligned} \mathbf{E}(a_n^{(3)})^2 &\sim \frac{1}{n^2} \sum_{m \geq 0} C_m^2 (2m)! \sum_{i,j \geq 1; i \neq j}^{n-1} R_{H_1} \left(\frac{i}{n}, \frac{j}{n} \right)^{2m} \left(\frac{i}{n} \right)^{-2H_1(m+\frac{1}{2})} \left(\frac{j}{n} \right)^{-2H_1(m+\frac{1}{2})} d_{i,\varepsilon,n,m} d_{j,\varepsilon,n,m} \\ &:= \sum_{m \geq 0} C_m^2 (2m)! A_m(\varepsilon, n). \end{aligned}$$

where

$$A_m(\varepsilon, n) = \frac{1}{n^2} \sum_{i,j \geq 1; i \neq j}^{n-1} R_{H_1} \left(\frac{i}{n}, \frac{j}{n} \right)^{2m} \left(\frac{i}{n} \right)^{-2H_1(m+\frac{1}{2})} \left(\frac{j}{n} \right)^{-2H_1(m+\frac{1}{2})} d_{i,\varepsilon,n,m} d_{j,\varepsilon,n,m}.$$

We can now claim that, for every fixed $m \geq 1$

$$\lim_{\varepsilon \rightarrow 0} \lim_{n \rightarrow \infty} A_m(\varepsilon, n) = 0. \quad (24)$$

Indeed, for every $m \geq 0$, we get

$$\begin{aligned}
|d_{i,\varepsilon,n,m}| &= \left| \left(\left(\frac{\left(\frac{i}{n}\right)^{2H_1}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \varepsilon\right)} \right)^{m+\frac{1}{2}} - 1 + 1 - \left(\frac{\left(\frac{i}{n}\right)^{2H_1}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)}\right)} \right)^{m+\frac{1}{2}} \right) \right| \\
&\leq \left| 1 - \left(\frac{\left(\frac{i}{n}\right)^{2H_1}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \varepsilon\right)} \right)^{m+\frac{1}{2}} \right| + \left| 1 - \left(\frac{\left(\frac{i}{n}\right)^{2H_1}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)}\right)} \right)^{m+\frac{1}{2}} \right| \\
&\leq \left| 1 - \left(\frac{\left(\frac{i}{n}\right)^{2H_1}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \varepsilon\right)} \right)^{m+1} \right| + \left| 1 - \left(\frac{\left(\frac{i}{n}\right)^{2H_1}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)}\right)} \right)^{m+1} \right| \\
&= c_m \left(\left| \left(\frac{\varepsilon}{\left(\left(\frac{i}{n}\right)^{2H_1} + \varepsilon\right)} \right) \right| + \left| \left(\frac{n^{-2(\alpha+H_1)}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)}\right)} \right) \right| \right).
\end{aligned}$$

Now, for every i, n, m , we have $\lim_{\varepsilon \rightarrow 0} \left| \left(\frac{\varepsilon}{\left(\left(\frac{i}{n}\right)^{2H_1} + \varepsilon\right)} \right) \right| = 0$ and for every $i \geq 1$,

$$\left| \left(\frac{n^{-2(\alpha+H_1)}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)}\right)} \right) \right| \leq \left| \left(\frac{n^{-2(\alpha+H_1)}}{\left(\left(\frac{1}{n}\right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)}\right)} \right) \right| \leq c \frac{n^{2H_1}}{n^{2(\alpha+2H_1)} + n^{2H_1}} \xrightarrow{n \rightarrow +\infty} 0$$

because $\alpha > 0$.

Furthermore, we know that

$$\frac{1}{n^2} \sum_{i,j=0}^{n-1} R_{H_1} \left(\frac{i}{n}, \frac{j}{n} \right)^{2m} \left(\frac{i}{n} \right)^{-2H_1(m+\frac{1}{2})} \left(\frac{j}{n} \right)^{-2H_1(m+\frac{1}{2})}$$

converges as $n \rightarrow \infty$ to $\int_0^1 \int_0^1 R(u, v)^{2m} (uv)^{-2H_1(m+\frac{1}{2})} du dv$. Since this quantity is finite ([3] and [4]), it implies (24).

We will now prove that

$$\sum_{m \geq 0} C_m^2 (2m)! \sup_{n, \varepsilon} |A_m(\varepsilon, n)| < \infty. \quad (25)$$

Relation (24) and (25) will imply the convergence of $a_n^{(3)}$ to zero in $L^2(\Omega)$. We need to find an upper bound for the terms $|d_{i,\varepsilon,n,m}|$ and $|d_{j,\varepsilon,n,m}|$ in order to continue.

$$\begin{aligned} d_{i,\varepsilon,n,m} &= \left(\left(\frac{\left(\frac{i}{n}\right)^{2H_1}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \varepsilon\right)} \right)^{m+\frac{1}{2}} - \left(\frac{\left(\frac{i}{n}\right)^{2H_1}}{\left(\left(\frac{i}{n}\right)^{2H_1} + \frac{1}{2}n^{-2(\alpha+H_1)}\right)} \right)^{m+\frac{1}{2}} \right) \\ &= \left(\left(\frac{1}{\left(1 + \varepsilon n^{2H} i^{-2H}\right)} \right)^{m+\frac{1}{2}} - \left(\frac{1}{\left(1 + \frac{1}{2}n^{-2\alpha} i^{-2H}\right)} \right)^{m+\frac{1}{2}} \right). \end{aligned}$$

One can note that

$$0 \leq \left(\frac{1}{\left(1 + \varepsilon n^{2H} i^{-2H}\right)} \right)^{m+\frac{1}{2}} \leq 1 \quad \text{and} \quad 0 \leq \left(\frac{1}{\left(1 + \frac{1}{2}n^{-2\alpha} i^{-2H}\right)} \right)^{m+\frac{1}{2}} \leq 1$$

because $\varepsilon n^{2H} i^{-2H} > 0$. From the above inequalities, we can deduce that

$$-1 \leq \left(\frac{1}{\left(1 + \varepsilon n^{2H} i^{-2H}\right)} \right)^{m+\frac{1}{2}} - \left(\frac{1}{\left(1 + \frac{1}{2}n^{-2\alpha} i^{-2H}\right)} \right)^{m+\frac{1}{2}} \leq 1$$

and finally,

$$|d_{i,\varepsilon,n,m}| \leq 1 \quad \text{and} \quad |d_{j,\varepsilon,n,m}| \leq 1.$$

By bounding from above the terms $|d_{i,\varepsilon,n,m}|$ and $|d_{j,\varepsilon,n,m}|$ by 1 in $\sum_{m \geq 0} C_m^2 (2m)! \sup_{n,\varepsilon} |A_m(\varepsilon, n)|$ we obtain that

$$\begin{aligned} \sum_{m \geq 0} C_m^2 (2m)! \sup_{n,\varepsilon} |A_m(\varepsilon, n)| &\leq \sum_{m \geq 0} C_m^2 (2m)! \sup_n \frac{1}{n^2} \sum_{i,j \geq 1, i \neq j}^{n-1} R_{H_1} \left(\frac{i}{n}, \frac{j}{n} \right)^{2m} \left(\frac{i}{n} \right)^{-2H_1(m+\frac{1}{2})} \left(\frac{j}{n} \right)^{-2H_1(m+\frac{1}{2})} \\ &= \sum_{m \geq 0} C_m^2 (2m)! \sup_n \frac{1}{n^2} \sum_{i,j \geq 1, i \neq j}^{n-1} R_{H_1} \left(1, \left(\frac{j}{i} \right) \right)^{2m} \left(\frac{j}{i} \right)^{-2H_1 m} \left(\frac{i}{n} \frac{j}{n} \right)^{-H_1}. \end{aligned}$$

Let's focus on the case where $H_1 < \frac{1}{2}$ first. Let $Q_{H_1}(z)$ be the function defined by

$$Q_{H_1}(z) = \begin{cases} \frac{R_{H_1}(1,z)}{z^{H_1}} & \text{if } z \in (0, 1] \\ 0 & \text{if } z = 0. \end{cases}$$

For $H_1 < \frac{1}{2}$, we have

$$Q_{H_1}(z) \leq z^{H_1}.$$

Indeed, the function $f(z) = 1 - z^{2H_1} - (1 - z)^{2H_1}$ is negative on $[0, 1]$, increasing on $[\frac{1}{2}, 1]$, decreasing on $[0, \frac{1}{2}]$ and $f(1) = f(0) = 0$. It follows that

$$\begin{aligned}
\sum_{m \geq 0} C_m^2(2m)! \sup_{n, \varepsilon} |A_m(\varepsilon, n)| &\leq 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{1}{n^2} \sum_{i, j=0; i > j}^{n-1} \left(\frac{j}{i}\right)^{2H_1 m} \left(\frac{i}{n} \frac{j}{n}\right)^{-H_1} \\
&\leq 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{1}{n^2} \sum_{i, j=0; i > j}^{n-1} \left(\frac{j}{n}\right)^{H_1(2m-1)} \left(\frac{i}{n}\right)^{-H_1(2m+1)} \\
&\leq 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n n^{2H_1-2} \sum_{i=0}^{n-1} i^{-H_1(2m+1)} \sum_{j=1}^{i-1} \int_j^{j+1} j^{H_1(2m-1)} dx \\
&\leq 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n n^{2H_1-2} \sum_{i=0}^{n-1} i^{-H_1(2m+1)} \int_0^i x^{H_1(2m-1)} dx \\
&\leq 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{n^{2H_1-2}}{2H_1 m - H_1 + 1} \sum_{i=0}^{n-1} i^{1-2H_1} \\
&\leq 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{n^{-1}}{2H_1 m - H_1 + 1} \sum_{i=0}^{n-1} 1 \\
&\leq 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{1}{2H_1 m - H_1 + 1} \leq 2 \sum_{m \geq 0} \frac{C_m^2(2m)!}{2H_1 m - H_1 + 1}.
\end{aligned}$$

Given that, by using Stirling's formula, the coefficient $C_m^2(2m)!$ behaves as $\frac{1}{\sqrt{m}}$, we obtain that the above sum is finite. Thus, we obtain the convergence of $a_n^{(3)}$ to zero in $L^2(\Omega)$ for $H_1 < \frac{1}{2}$.

Let us now treat the case $H_1 > \frac{1}{2}$. We know (see [4], Lemma 1) that the function Q_H is increasing on $[0, 1]$. Since $\frac{j}{i} \leq \frac{i-1}{i} = 1 - \frac{1}{i}$ it holds that $Q_H(\frac{j}{i}) \leq Q_H(1 - \frac{1}{i})$. Then

$$\begin{aligned}
\sum_{m \geq 0} C_m^2(2m)! \sup_{n, \varepsilon} |A_m(\varepsilon, n)| &\leq 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{1}{n^2} \sum_{i=1}^{n-1} Q_H\left(1 - \frac{1}{i}\right) \sum_{j=1}^{i-1} \left(\frac{i}{n} \frac{j}{n}\right)^{-H_1} \\
&= 2 \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{1}{n} \sum_{i=1}^{n-1} Q_H\left(1 - \frac{1}{i}\right) \left(\frac{i}{n}\right)^{-H_1} \sum_{j=1}^{i-1} \int_{\frac{j-1}{n}}^{\frac{j}{n}} x^{-H_1} dx \\
&\leq c_H \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{1}{n} \sum_{i=1}^{n-1} Q_H\left(1 - \frac{1}{i}\right) \left(\frac{i}{n}\right)^{-H_1} \left(\frac{i-1}{n}\right)^{1-H_1} \\
&\sim c_H \sum_{m \geq 0} C_m^2(2m)! \sup_n \frac{1}{n} \sum_{i=1}^{n-1} Q_H\left(1 - \frac{1}{i}\right) \left(\frac{i}{n}\right)^{1-2H_1}.
\end{aligned}$$

By adapting Lemma 2 in [4] (by separating the sum over i in a sum with $\frac{1}{i} \leq \delta$ and $\frac{1}{i} > \delta$ with δ suitably chosen), we can prove that

$$\frac{1}{n} \sum_{i,j=0}^{n-1} R_{H_1} \left(1, \left(\frac{j}{i} \right) \right)^{2m} \left(\frac{j}{i} \right)^{-2H_1 m} \left(\frac{i}{n} \frac{j}{n} \right)^{-H} \leq c(H_1) m^{-\frac{1}{2H_1}}$$

with $c(H_1)$ not depending on m nor n . As a consequence

$$\sum_{m \geq 0} c_m^2 (2m)! \sup_{n, \varepsilon} |A_m(\varepsilon, n)| \leq c(H_1) c_m^2 (2m)! m^{-\frac{1}{2H_1}}.$$

The Stirling formula implies again that the above series is finite. ■

Theorem 1 *Let $\langle S \rangle_n$ be given by (3). Then, as $n \rightarrow \infty$, we have the convergence in distribution*

$$n^{\alpha+H_1-1} \langle S \rangle_n \rightarrow \int_{\mathbb{R}} K^2(y) dy L^{H_1}(1, 0)$$

where $L^{H_1}(1, 0)$ is the local time of the fractional Brownian motion B^{H_1} .

Proof: Using Proposition 2 it suffices to check that $n^{\alpha+H_1} \int_0^1 K^2(n^{\alpha+H_1} B_s^{H_1}) ds$ converges to $\int_{\mathbb{R}} K^2(y) dy L^{H_1}(1, 0)$. Using the occupation time formula (17), we obtain

$$n^{\alpha+H_1} \int_0^1 K^2(n^{\alpha+H_1} B_s^{H_1}) ds = n^{\alpha+H_1} \int_{\mathbb{R}} K^2(n^{\alpha+H_1} x) L^{H_1}(1, x) dx = \int_{\mathbb{R}} K^2(y) L(1, y n^{-\alpha-H_1}) dy$$

which converges as $n \rightarrow \infty$ to $\int_{\mathbb{R}} K^2(y) dy L^{H_1}(1, 0)$ by using the continuity properties of the local time. ■

5 Limit distribution of S_n

In this paragraph, we prove the limit in distribution of (7). Recall the notation (13) and let's consider the Gaussian vector

$$X^{H_2} = (X_1^{H_2}, \dots, X_n^{H_2}) = (B_1^{H_2} - B_0^{H_2}, \dots, B_n^{H_2} - B_{n-1}^{H_2}).$$

From this definition, it follows that

$$S_n = \sum_{i=0}^{n-1} K(n^{\alpha} B_i^{H_1}) (B_{i+1}^{H_2} - B_i^{H_2}) = \sum_{i=0}^{n-1} K(n^{\alpha} B_i^{H_1}) X_{i+1}^{H_2}.$$

Theorem 2 Let (S_n) be given by (7) and assume that

$$\alpha < 1 - H_1 \quad (26)$$

Then we have the convergence in law

$$n^{\alpha+H_1-1} S_n \xrightarrow{n \rightarrow +\infty} d_1 W_{L^{H_1}(1,0)}$$

where $L^{H_1}(1,0)$ is the local time of B^{H_1} , $d_1 := \int_{\mathbb{R}} K^2(y) dy$ and W is a Brownian motion independent from B^{H_1} .

Proof: We will study the characteristic function of $n^{\frac{\alpha}{2} + \frac{H_1}{2} - \frac{1}{2}} S_n$. In order to simplify the presentation, we will use the following notation. Let i_0 be the imaginary unit and λ_n be

$$\lambda_n = \lambda n^{\frac{\alpha}{2} + \frac{H_1}{2} - \frac{1}{2}} \quad \text{with } \lambda \in \mathbb{R}.$$

Using the independence of the two fBms and computing the conditional expectation of $e^{i\lambda_n S_n}$ given B^{H_1} we get

$$\mathbf{E} \left(e^{i_0 \lambda_n S_n} \right) = \mathbf{E} \left(e^{-\frac{1}{2} \sum_{i,j=0}^{n-1} \lambda_n^2 K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) f_{H_2}(i,j)} \right)$$

because if X is a Gaussian vector with mean μ and covariance matrix Σ , it's characteristic function is given by

$$\mathbf{E} \left(e^{i_0 \langle t, X \rangle} \right) = e^{i_0 \mu^T t - \frac{1}{2} t^T \Sigma t}.$$

It follows that, with $f_{H_2}(i,j)$ given by (13),

$$\begin{aligned} \mathbf{E} \left(e^{i_0 \lambda_n S_n} \right) &= \mathbf{E} \left(e^{-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})} e^{-\frac{\lambda_n^2}{2} \sum_{i \neq j=0}^{n-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) f_{H_2}(i,j)} \right) \\ &= \mathbf{E} \left(e^{-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})} e^{-\lambda_n^2 \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) f_{H_2}(i,j)} \right) \\ &= \mathbf{E} \left(e^{-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})} e^{-\lambda_n^2 \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) H_2(2H_2-1) \int_i^{i+1} \int_j^{j+1} |s-u|^{2H_2-2} duds} \right) \\ &= \mathbf{E} \left(e^{-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})} e^{-\lambda_n^2 H_2(2H_2-1) \int_0^n \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) |s-u|^{2H_2-2} duds} \right). \end{aligned}$$

Consider the process $(V_n)_{n \geq 0}$ defined by

$$V_n = \int_0^n \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) |s-u|^{2H_2-2} duds$$

and the function ψ defined by $\psi(x) = e^{-\lambda_n^2 H_2(2H_2-1)x}$. Note that, since we excluded the diagonal, the integral $duds$ in the expression of V_n makes sense even for $H_2 < \frac{1}{2}$. Note also that V_n is a bounded variation process (its quadratic variation is 0). Furthermore,

$$\psi'(x) = -\lambda_n^2 H_2(2H_2-1) e^{-\lambda_n^2 H_2(2H_2-1)x}.$$

Using the change of variables formula for bounded variation processes, it follows that

$$\psi(V_n) = 1 + \int_0^n \psi'(V_s) dV_s$$

i.e.,

$$e^{-\lambda_n^2 H_2(2H_2-1)V_n} = 1 - \lambda_n^2 H_2(2H_2-1) \int_0^n e^{-\lambda_n^2 H_2(2H_2-1)V_s} dV_s.$$

Therefore,

$$\begin{aligned} \mathbf{E} \left(e^{i_0 \lambda_n S_n} \right) &= \mathbf{E} \left(e^{-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})} \left(1 - \lambda_n^2 H_2(2H_2-1) \int_0^n e^{-\lambda_n^2 H_2(2H_2-1)V_s} dV_s \right) \right) \\ &= \mathbf{E} \left(e^{-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})} \right) \\ &\quad - \mathbf{E} \left(\lambda_n^2 H_2(2H_2-1) e^{-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})} \int_0^n e^{-\lambda_n^2 H_2(2H_2-1)V_s} dV_s \right) \\ &:= \mathbf{E}(T_1) - \mathbf{E}(T_2). \end{aligned}$$

We will now focus on the term $\mathbf{E}(T_2)$ and show that

$$T_2 \xrightarrow{L^1} 0.$$

From

$$dV_s = \left(\int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) |s-u|^{2H_2-2} du \right) ds$$

we get

$$\begin{aligned} \mathbf{E}(T_2) &= \mathbf{E} \left(\lambda_n^2 H_2(2H_2-1) e^{-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})} \right. \\ &\quad \times \left. \int_0^n e^{-\lambda_n^2 H_2(2H_2-1)V_s} \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) |s-u|^{2H_2-2} duds \right) \\ &= \mathbf{E} \left(\lambda_n^2 H_2(2H_2-1) \int_0^n e^{-\frac{\lambda_n^2}{2} \int_0^s K^2(n^\alpha B_{[u]}^{H_1}) du} e^{-\frac{\lambda_n^2}{2} \int_s^n K^2(n^\alpha B_{[u]}^{H_1}) du} \right. \\ &\quad \times \left. e^{-\lambda_n^2 H_2(2H_2-1)V_s} \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) |s-u|^{2H_2-2} duds \right). \end{aligned}$$

Recall that the following holds

$$\mathbf{E} \left(e^{i_0 \lambda_n S_s} | B_s^{H_1} \right) = \mathbf{E} \left(e^{-\frac{\lambda_n^2}{2} \int_0^s K^2(n^\alpha B_{[u]}^{H_1}) du} e^{-\lambda_n^2 H_2(2H_2-1)V_s} | B_s^{H_1} \right). \quad (27)$$

This can be seen for s integer as at the beginning of this proof and also (27) can easily be checked for any $s > 0$. We will use this property to compute the following upper bound for $\mathbf{E}(|T_2|)$

$$\begin{aligned}
\mathbf{E}(|T_2|) &\leq \mathbf{E} \left(\lambda_n^2 \int_0^n e^{-\frac{\lambda_n^2}{2} \int_0^s K^2(n^\alpha B_{[u]}^{H_1}) du} e^{-\lambda_n^2 H_2 (2H_2 - 1) V_s} \underbrace{\left| e^{-\frac{\lambda_n^2}{2} \int_s^n K^2(n^\alpha B_{[u]}^{H_1}) du} \right|}_{\leq 1} \right. \\
&\quad \times \left. \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) H_2 |2H_2 - 1| |s - u|^{2H_2 - 2} duds \right) \\
&\leq \mathbf{E} \left(\lambda_n^2 \int_0^n \mathbf{E} \left(e^{-\frac{\lambda_n^2}{2} \int_0^s K^2(n^\alpha B_{[u]}^{H_1}) du} e^{-\lambda_n^2 H_2 (2H_2 - 1) V_s} \middle| B_s^{H_1} \right) \right. \\
&\quad \times \left. \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) H_2 |2H_2 - 1| |s - u|^{2H_2 - 2} duds \right).
\end{aligned}$$

This is true because all the terms of the double integral are measurable with respect to the filtration generated by $(B_u^{H_1}, u \leq s)$. At this point, we use (27) to write

$$\begin{aligned}
\mathbf{E}(|T_2|) &\leq \mathbf{E} \left(\lambda_n^2 \int_0^n \mathbf{E} \left(e^{i_0 \lambda_n S_s} \middle| B_s^{H_1} \right) \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) H_2 |2H_2 - 1| |s - u|^{2H_2 - 2} duds \right) \\
&\leq \mathbf{E} \left(\lambda_n^2 \int_0^n \underbrace{|e^{i_0 \lambda_n S_s}|}_{=1} \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) H_2 |2H_2 - 1| |s - u|^{2H_2 - 2} duds \right) \\
&\leq \mathbf{E} \left(\lambda_n^2 \int_0^n \int_0^{[s]} K(n^\alpha B_{[s]}^{H_1}) K(n^\alpha B_{[u]}^{H_1}) H_2 |2H_2 - 1| |s - u|^{2H_2 - 2} duds \right) \\
&\leq \mathbf{E} \left(\lambda_n^2 \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) H_2 |2H_2 - 1| \int_i^{i+1} \int_j^{j+1} |s - u|^{2H_2 - 2} duds \right).
\end{aligned}$$

Assume that $H_2 > \frac{1}{2}$, ergo $|2H_2 - 1| > 0$ and $f_{H_2}(i, j) > 0$. Consequently,

$$\begin{aligned}
\mathbf{E}(|T_2|) &\leq \mathbf{E} \left(\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) f_{H_2}(i, j) \right) \\
&\leq \mathbf{E} \left(\frac{\lambda_n^2}{2} n^{\alpha + H_1 - 1} \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) f_{H_2}(i, j) \right).
\end{aligned}$$

The previous term is exactly the non-diagonal term of the L^2 -norm of $n^{\frac{\alpha}{2} + \frac{H_1}{2} - \frac{1}{2}} S_n$ and we know that under condition (11), it converges to zero when $n \rightarrow +\infty$. Finally we have

$$\mathbf{E}(|T_2|) \xrightarrow{n \rightarrow +\infty} 0.$$

Assume now that $H_2 < \frac{1}{2}$. It follows that $|2H_2 - 1| < 0$ and $f_{H_2}(i, j) < 0$, which gives us

$$\begin{aligned} \mathbf{E}(|T_2|) &\leq \mathbf{E} \left(-\frac{\lambda_n^2}{2} \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) f_{H_2}(i, j) \right) \\ &\leq \mathbf{E} \left(-\frac{\lambda^2}{2} n^{\alpha+H_1-1} \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) f_{H_2}(i, j) \right). \end{aligned}$$

As in the previous case, this term is again exactly the non-diagonal term of the L^2 -norm of $n^{\frac{\alpha}{2} + \frac{H_1}{2} - \frac{1}{2}} S_n$ and for the same reasons, we get the following result again (which is now valid for any $H_2 \in (0, 1)$)

$$\mathbf{E}(|T_2|) \xrightarrow{n \rightarrow +\infty} 0.$$

Concerning the term T_1 , we note that

$$\mathbf{E}(T_1) = \mathbf{E} \left(e^{-\frac{\lambda^2}{2} \langle S \rangle_n} \right)$$

and the result follows from Theorem 1. ■

Remark 1 *The following comments deal with the conditions (11) and (26). Condition (26) is a natural extension of the condition $\alpha < \frac{1}{2}$ in e.g. [15], [16] which means that the bandwidth parameter satisfies $nh_n^2 = nn^{-2\alpha} \rightarrow \infty$ as $n \rightarrow \infty$. From (11) and (26), this is the constraint we find for α (considering α is our degree of freedom)*

$$\left\{ \begin{array}{l} 0 < H_1 < 1 \\ 0 < H_2 < 1 \\ \alpha > 4H_2 - H_1 - 2 \\ \alpha < 1 - H_1 \end{array} \right\} \Leftrightarrow \left\{ \begin{array}{l} 0 < H_1 < 1 \\ 0 < H_2 < 1 \\ 4H_2 - H_1 - 2 < \alpha < 1 - H_1. \end{array} \right.$$

As an example, consider the case where $H_1 = H_2 = H$. Those constraints become

$$\left\{ \begin{array}{l} 0 < H < 1 \\ (3H - 2)^+ < \alpha < 1 - H. \end{array} \right.$$

For this system to have a solution, we need to verify that

$$3H - 2 < 1 - H \Leftrightarrow H < \frac{3}{4}.$$

As a result, our constraints become

$$\begin{cases} 0 < H < \frac{3}{4} \\ (3H - 2)^+ < \alpha < 1 - H. \end{cases}$$

We could also consider the case where α has a fixed value and where the constraints would be on H_1 and H_2 .

6 The stable convergence

In this section we will study the convergence of the vector $(S_n, (G_t)_{t \geq 0})$ where $(G_t)_{t \geq 0}$ is a stochastic process independent from B^{H_1} and satisfies some additional conditions. In this case, since the process $(G_t)_{t \geq 0}$ is not necessarily a Gaussian process and since no information is available on the correlation between B^{H_2} and G_t , the characteristic function of the vector $(S_n, (G_t)_{t \geq 0})$ cannot be computed directly. To compute it, we will use the tools of the stochastic calculus with respect to the fractional Brownian motion. The basic observation is that S_n can be expressed as a stochastic integral with respect to B^{H_2} . Indeed,

$$\begin{aligned} S_n &= \sum_{i=0}^{n-1} K(n^\alpha B_i^{H_1})(B_{i+1}^{H_2} - B_i^{H_2}) = \sum_{i=0}^{n-1} K(n^\alpha B_i^{H_1})\delta^{H_2}(\mathbf{1}_{[i, i+1]}(\cdot)) \\ &= \sum_{i=0}^{n-1} \delta^{H_2}(K(n^\alpha B_i^{H_1})\mathbf{1}_{[i, i+1]}(\cdot)) + \left\langle \underbrace{D^{H_2} K(n^\alpha B_i^{H_1})}_{=0 \text{ from } B^{H_1} \perp B_t^{H_2}}, \mathbf{1}_{[i, i+1]}(\cdot) \right\rangle_{\mathcal{H}_{H_2}} \\ &= \sum_{i=0}^{n-1} \int_i^{i+1} K(n^\alpha B_i^{H_1}) dB_s^{H_2} = \sum_{i=0}^{n-1} \int_i^{i+1} K(n^\alpha B_{[s]}^{H_1}) dB_s^{H_2} = \int_0^n K(n^\alpha B_{[s]}^{H_1}) dB_s^{H_2}. \end{aligned} \tag{28}$$

We will also use the “bracket” of S_n . This quantity equals

$$\begin{aligned} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) &= \sum_{i=0}^{n-1} \int_i^{i+1} K^2(n^\alpha B_i^{H_1}) ds \\ &= \sum_{i=0}^{n-1} \int_i^{i+1} K^2(n^\alpha B_{[s]}^{H_1}) ds = \int_0^n K^2(n^\alpha B_{[s]}^{H_1}) ds. \end{aligned}$$

Before going any further, we will describe the elements of the stochastic calculus with respect to fractional Brownian motion that we will be using in the sequel. We will start by introducing some notations and definitions. Let ϕ be the function defined by

$$\phi(s, t) = H(2H - 1)|s - t|^{2H-2}, \quad s, t \in \mathbb{R}.$$

Let D (introduced in section 2) be the Malliavin derivative operator with respect to the fractional Brownian motion with Hurst parameter H . Based on this operator, let D^ϕ be another derivative operator (called the ϕ -derivative operator) defined by

$$D_t^\phi F = \int_{\mathbb{R}} \phi(t, v) D_v F dv$$

for any F in the domain of D . For more details about this operator, see [2]. Let $\mathcal{L}_\phi(0, T)$ be the family of stochastic processes F on $[0, T]$ with the following properties: $F \in \mathcal{L}_\phi(0, T)$ if and only if $\mathbf{E} \left[\|F\|_{\mathcal{H}}^2 \right] < \infty$, F is ϕ -differentiable, the trace of $D_s^\phi F_t$, $0 \leq s, t \leq T$, exists, and $\mathbf{E} \left[\int_0^T \int_0^T \left| D_s^\phi F_t \right|^2 ds dt \right] < \infty$ and for each sequence of partitions $(\pi_n, n \in \mathbb{N})$ such that $|\pi_n| \rightarrow 0$ as $n \rightarrow +\infty$,

$$\sum_{i,j=0}^{n-1} \mathbf{E} \left[\int_{t_i^{(n)}}^{t_{i+1}^{(n)}} \int_{t_j^{(n)}}^{t_{j+1}^{(n)}} \left| D_s^\phi F_{t_i^{(n)}}^\pi D_t^\phi F_{t_j^{(n)}}^\pi - D_s^\phi F_t D_t^\phi F_s \right| ds dt \right]$$

and

$$\mathbf{E} \left[\|F^\pi - F\|_{\mathcal{H}}^2 \right]$$

tend to 0 as $n \rightarrow +\infty$, where $\pi_n := 0 = t_0^{(n)} < t_1^{(n)} < \dots < t_{n-1}^{(n)} < t_n^{(n)} = T$.

In our particular situation, we are dealing with processes of the form $\int_0^t F_u dB_u^H + \int_0^t G_u du$, (where B^H is a fractional Brownian motion with Hurst parameter H), for which the following Itô formula holds in the case $H > \frac{1}{2}$.

Theorem 3 Let $\eta_t = \int_0^t F_u dB_u^H + \int_0^t G_u du$, for $t \in [0, T]$ with $\mathbf{E} \left[\sup_{0 \leq s \leq T} |G_s| \right] < \infty$ and let $(F_u, 0 \leq u \leq T)$ be a stochastic process in $\mathcal{L}_\phi(0, T)$. Assume that there is a $\beta > 1 - H$ such that

$$\mathbf{E} \left[|F_u - F_v|^2 \right] \leq C |u - v|^{2\beta} \quad (29)$$

where $|u - v| \leq \zeta$ for some $\zeta > 0$ and

$$\lim_{0 \leq u, v \leq t, |u-v| \rightarrow 0} \mathbf{E} \left[\left| D_u^\phi(F_u - F_v) \right|^2 \right] = 0. \quad (30)$$

Let $f : \mathbb{R}_+ \times \mathbb{R} \rightarrow \mathbb{R}$ be a function having the first continuous derivative in its first variable and the second continuous derivative in its second variable. Assume that these derivatives are bounded. Moreover, it is assumed that $\mathbf{E} \left[\int_0^T \left| F_s D_s^\phi \eta_s \right| ds \right] < \infty$ and $(\frac{\partial f(s, \eta_s)}{\partial x} F_s, s \in [0, T]) \in \mathcal{L}_\phi(0, T)$. Then for $t \in [0, T]$,

$$\begin{aligned} f(t, \eta_t) &= f(0, 0) + \int_0^t \frac{\partial f}{\partial s}(s, \eta_s) ds + \int_0^t \frac{\partial f}{\partial x}(s, \eta_s) G_s ds \\ &\quad + \int_0^t \frac{\partial f}{\partial x}(s, \eta_s) F_s dB_s^H + \int_0^t \frac{\partial^2 f}{\partial x^2}(s, \eta_s) F_s D_s^\phi \eta_s ds. \end{aligned} \quad (31)$$

We also have the following technical lemma ([2] p.71) that will be particularly useful for our future computations.

Lemma 4 *Let $(F_t, t \in [0, T])$ be a stochastic process in $\mathcal{L}_\phi(0, T)$ and*

$$\sup_{0 \leq s \leq T} \mathbf{E} \left[\left| D_s^\phi F_s \right|^2 \right] < \infty$$

and let $\eta_t = \int_0^t F_u dB_u^H$ for $t \in [0, T]$. Then, for $s, t \in [0, T]$,

$$D_s^\phi \eta_t = \int_0^t D_s^\phi F_u dB_u^H + \int_0^t F_u \phi(s, u) du. \quad (32)$$

It is now possible to state the main result of this section

Theorem 4 *Assume that (11) and (26) holds. Let $(G_t)_{t \geq 0}$ be a stochastic process independent from B^{H_1} and adapted to the filtration generated by B^{H_2} such that for every $t \geq 0$ the random variable G_t belongs to $\mathbb{D}^{1,2}$ and $\|D_s G_t\| \leq C$ for any s, t and ω . Then the vector $(S_n, (G_t)_{t \geq 0})$ converges in the sense of finite dimensional distributions to the vector $(cW_{L^{H_1}(1,0)}, (G_t)_{t \geq 0})$, where c is a positive constant.*

Proof: In order to simplify the presentation, the following notations will be used. We will denote by λ_n (like we did in a previous proof) the quantity

$$\lambda_n = \lambda n^{\frac{\alpha}{2} + \frac{H_1}{2} - \frac{1}{2}}$$

where $\lambda \in \mathbb{R}$. The following notation will also be used:

$$e(\lambda, n) = e^{-\frac{\lambda^2}{2} n^{\alpha+H_1-1} \int_0^n K^2(n^\alpha B_{[u]}^{H_1}) du} = e^{-\frac{\lambda^2}{2} n^{\alpha+H_1-1} \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1})}.$$

Let $(F_t, t \geq 0)$ and $(G_t, t \geq 0)$ be two stochastic processes defined by

$$\begin{cases} F_u = K(n^\alpha B_{[u]}^{H_1}) \\ G_u = -i_0 \frac{\lambda_n}{2} K^2(n^\alpha B_{[u]}^{H_1}) \end{cases}$$

and let $(\eta_t^{(\lambda_n)}, t \geq 0)$ be the stochastic process defined by

$$\eta_t^{(\lambda_n)} = \int_0^t F_u dB_u^H + \int_0^t G_u du = \int_0^t K(n^\alpha B_{[u]}^{H_1}) dB_u^{H_2} - i_0 \frac{\lambda_n}{2} \int_0^t K^2(n^\alpha B_{[u]}^{H_1}) du.$$

Consider the function $f : \mathbb{C} \rightarrow \mathbb{C}, f(x) = e^{i_0 \lambda_n x}$. We can apply the Itô formula to $f(\eta_t^{(\lambda_n)})$ in order to obtain

$$\begin{aligned} e^{i_0 \lambda_n \eta_t^{(\lambda_n)}} &= 1 + \frac{\lambda_n^2}{2} \int_0^t e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K^2(n^\alpha B_{[s]}^{H_1}) ds \\ &\quad + i_0 \lambda_n \int_0^t e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_{[s]}^{H_1}) dB_s^{H_2} \\ &\quad - \lambda_n^2 \int_0^t e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_{[s]}^{H_1}) D_s^{\phi, H_2} \eta_s^{(\lambda_n)} ds \end{aligned} \quad (33)$$

where D^{ϕ, H_2} is the operator D^ϕ introduced above with respect to the fractional Brownian motion B^{H_2} . We use Lemma 4 to compute $D_s^{\phi, H_2} \eta_s^{(\lambda_n)}$. We get

$$\begin{aligned}
D_s^{\phi, H_2} \eta_s^{(\lambda_n)} &= D_s^{\phi, H_2} \int_0^s K(n^\alpha B_{[u]}^{H_1}) dB_u^{H_2} - i_0 \frac{\lambda_n}{2} \underbrace{D_s^{\phi, H_2} \int_0^s K^2(n^\alpha B_{[u]}^{H_1}) du}_{=0 \text{ from } B^{H_1} \perp B_t^{H_2}} \\
&= \underbrace{\int_0^s D_s^{\phi, H_2} K(n^\alpha B_{[u]}^{H_1}) dB_u^{H_2}}_{=0 \text{ from } B^{H_1} \perp B_t^{H_2}} + H_2(2H_2 - 1) \int_0^s K(n^\alpha B_{[u]}^{H_1}) |s - u|^{2H_2-2} du \\
&= H_2(2H_2 - 1) \int_0^s K(n^\alpha B_{[u]}^{H_1}) |s - u|^{2H_2-2} du.
\end{aligned}$$

By substituting in (33), we obtain

$$\begin{aligned}
e^{i_0 \lambda_n \eta_t^{(\lambda_n)}} &= 1 + \frac{\lambda_n^2}{2} \int_0^t e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K^2(n^\alpha B_{[s]}^{H_1}) ds \\
&\quad + i_0 \lambda_n \int_0^t e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_{[s]}^{H_1}) dB_s^{H_2} \\
&\quad - \lambda_n^2 H_2(2H_2 - 1) \int_0^t e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_{[s]}^{H_1}) \\
&\quad \times \int_0^s K(n^\alpha B_{[u]}^{H_1}) |s - u|^{2H_2-2} duds. \tag{34}
\end{aligned}$$

By multiplying both sides of (34) by $e^{-\frac{\lambda_n^2}{2} \int_0^n K^2(n^\alpha B_{[u]}^{H_1}) du} = e(\lambda, n)$, we obtain

$$\begin{aligned}
e^{i_0 \lambda_n \int_0^n K(n^\alpha B_{[u]}^{H_1}) dB_u^{H_2}} &= e^{-\frac{\lambda_n^2}{2} \int_0^n K^2(n^\alpha B_{[u]}^{H_1}) du} \\
&\quad + \frac{\lambda_n^2}{2} \int_0^n e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K^2(n^\alpha B_{[s]}^{H_1}) ds \cdot e(\lambda, n) \\
&\quad + i_0 \lambda_n \int_0^n e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_{[s]}^{H_1}) dB_s^{H_2} \cdot e(\lambda, n) \\
&\quad - \lambda_n^2 H_2(2H_2 - 1) \int_0^n e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_{[s]}^{H_1}) \\
&\quad \times \int_0^s K(n^\alpha B_{[u]}^{H_1}) |s - u|^{2H_2-2} duds \cdot e(\lambda, n). \tag{35}
\end{aligned}$$

The sum of the last two terms in (35) can be written in a more suitable way by using sums instead of integrals. Together, these two last terms give us

$$\begin{aligned}
& \mathbf{E} \left(\lambda_n^2 \int_0^n e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K \left(n^\alpha B_{[s]}^{H_1} \right) \right. \\
& \quad \times \left[\frac{1}{2} K \left(n^\alpha B_{[s]}^{H_1} \right) - H_2(2H_2 - 1) \int_0^s K \left(n^\alpha B_{[u]}^{H_1} \right) |s - u|^{2H_2-2} du \right] ds \cdot e(\lambda, n) \Big) \\
= & \mathbf{E} \left(\lambda_n^2 \int_0^n e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K \left(n^\alpha B_{[s]}^{H_1} \right) \right. \\
& \quad \times \left[\frac{1}{2} K \left(n^\alpha B_{[s]}^{H_1} \right) - H_2(2H_2 - 1) \left(\int_0^i K \left(n^\alpha B_{[u]}^{H_1} \right) |s - u|^{2H_2-2} du \right. \right. \\
& \quad \left. \left. + \int_i^s K \left(n^\alpha B_i^{H_1} \right) |s - u|^{2H_2-2} du \right) \right] ds \cdot e(\lambda, n) \Big) \\
= & -\mathbf{E} \left(\lambda_n^2 \sum_{i=0}^{n-1} \int_i^{i+1} e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K \left(n^\alpha B_i^{H_1} \right) \right. \\
& \quad \times H_2(2H_2 - 1) \sum_{j=0}^{i-1} K \left(n^\alpha B_j^{H_1} \right) \int_j^{j+1} |s - u|^{2H_2-2} du ds \cdot e(\lambda, n) \Big) \\
& + \mathbf{E} \left(\lambda_n^2 \sum_{i=0}^{n-1} \int_i^{i+1} e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K^2 \left(n^\alpha B_i^{H_1} \right) \right. \\
& \quad \times \left[\frac{1}{2} - H_2(2H_2 - 1) \int_i^s |s - u|^{2H_2-2} du \right] ds \cdot e(\lambda, n) \Big). \tag{36}
\end{aligned}$$

Let us now fix $\beta_1, \dots, \beta_N \in \mathbb{R}$ and $t_1, \dots, t_N \geq 0$. We need to show that $\mathbf{E} \left(e^{i_0 \lambda_n S_n} e^{i_0(\beta_1 G_{t_1} + \dots + \beta_N G_{t_N})} \right)$ converges to $\mathbf{E} \left(e^{-\frac{\lambda^2(L^{H_1(1,0)})^2}{2}} e^{i_0(\beta_1 G_{t_1} + \dots + \beta_N G_{t_N})} \right)$. We will use the notation

$$g_N := e^{i_0(\beta_1 G_{t_1} + \dots + \beta_N G_{t_N})}.$$

By combining relations (35) and (36), we can write

$$\begin{aligned}
\mathbf{E} \left(e^{i_0 \lambda_n S_n} g_N \right) &= \mathbf{E}(e(\lambda, n) g_N) \\
&+ \mathbf{E} \left(i_0 \lambda_n \int_0^n e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_{[s]}^{H_1}) dB_s^{H_2} \cdot e(\lambda, n) g_N \right) \\
&- \mathbf{E} \left(\lambda_n^2 \sum_{i=0}^{n-1} \int_i^{i+1} e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_i^{H_1}) \right. \\
&\quad \times H_2(2H_2 - 1) \sum_{j=0}^{i-1} K(n^\alpha B_j^{H_1}) \int_j^{j+1} |s - u|^{2H_2-2} du ds \times e(\lambda, n) g_N \left. \right) \\
&+ \mathbf{E} \left(\lambda_n^2 \sum_{i=0}^{n-1} \int_i^{i+1} e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K^2(n^\alpha B_i^{H_1}) \right. \\
&\quad \times \left[\frac{1}{2} - H_2(2H_2 - 1) \int_i^s |s - u|^{2H_2-2} du \right] ds \times e(\lambda, n) g_N \left. \right) \\
&:= \mathbf{E}(e(\lambda, n) g_N) + T_1^* + T_2^* + T_3^*. \tag{37}
\end{aligned}$$

Let us begin by proving that the term T_2^* converges to zero as $n \rightarrow \infty$. Since

$$\left| e^{i_0 \lambda_n \eta_u^{(\lambda_n)}} \right| e^{-\frac{\lambda_n^2}{2} \int_0^n K^2(n^\alpha B_{[u]}^{H_1}) du} \leq 1 \tag{38}$$

for every $s \leq n$ and since $|e^{i_0 x}| = 1$ for every x real, T_2^* can be bounded as follows

$$\begin{aligned}
T_2^* &\leq \mathbf{E} \left(\lambda_n^2 c(H_2) \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \int_i^{i+1} \int_j^{j+1} |s - u|^{2H_2-2} du ds \right) \\
&\leq \mathbf{E} \left(\lambda^2 n^{\alpha+H_1-1} c(H_2) \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K(n^\alpha B_i^{H_1}) K(n^\alpha B_j^{H_1}) \int_i^{i+1} \int_j^{j+1} |s - u|^{2H_2-2} du ds \right)
\end{aligned}$$

and this goes to zero as in the proof showing that the non-diagonal term goes to zero under the renormalization $n^{\alpha+H_1-1}$. Let us now handle the term T_1^* . By using the independence of B^{H_1} and B^{H_2} we can write

$$T_1^* = \mathbf{E} \left(i_0 \lambda_n \int_0^n e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} K(n^\alpha B_{[s]}^{H_1}) e(\lambda, n) dB_s^{H_2} \cdot g_N \right).$$

The duality formula is used to obtain

$$\begin{aligned}
T_1^* &= \mathbf{E} \left(i_0 \lambda_n \langle \mathbf{1}_{[0,n]} e^{i_0 \lambda_n \eta \cdot (\lambda_n)} K(n^\alpha B_{[\cdot]}^{H_1}) e(\lambda, n), D^{H_2} g_N \rangle_{\mathcal{H}_{H_2}} \right) \\
&= \mathbf{E} \left(-\lambda_n \langle \mathbf{1}_{[0,n]} e^{i_0 \lambda_n \eta \cdot (\lambda_n)} K(n^\alpha B_{[\cdot]}^{H_1}) e(\lambda, n), g_N \sum_{k=1}^N \beta_k D^{H_2} G_{t_k} \rangle_{\mathcal{H}_{H_2}} \right).
\end{aligned}$$

Recall that the following formula holds (see [10] for further details)

$$\langle \phi, \psi \rangle_{\mathcal{H}_{H_2}} = H_2(2H_2 - 1) \int_0^T \int_0^T |r - u|^{2H_2-2} \phi_r \psi_u du dr$$

for any pair of functions in the Hilbert space \mathcal{H}_{H_2} . This formula is used to write T_1^* as

$$T_1^* = \mathbf{E} \left(-\lambda_n \sum_{k=1}^N \beta_k H_2(2H_2 - 1) \int_0^n \int_0^{t_k} e^{i_0 \lambda_n \eta_u} K(n^\alpha B_{[u]}^{H_1}) e(\lambda, n) D_v G_{t_k} |u - v|^{2H_2-2} dv du \right)$$

where the fact that G_t is adapted to the filtration of B^{H_2} is used. It suffices to show that for every fixed $t \geq 0$,

$$\lambda_n \mathbf{E} \left(\int_0^n \int_0^t e^{i_0 \lambda_n \eta_u} K(n^\alpha B_{[u]}^{H_1}) e(\lambda, n) D_v G_t |u - v|^{2H_2-2} dv du \right)$$

converges to zero as $n \rightarrow \infty$. Since the derivative of G_t is bounded and using (38) we find that the above term is less than

$$\begin{aligned} & \lambda_n c_1 \mathbf{E} \left(\int_0^n \int_0^t K(n^\alpha B_{[u]}^{H_1}) |u - v|^{2H_2-2} dv du \right) \\ &= \lambda_n c_1 \sum_{i=0}^{n-1} \mathbf{E} \left(K(n^\alpha B_i^{H_1}) \int_i^{i+1} \int_0^t |u - v|^{2H_2-2} dv du \right) \\ &= \lambda_n c_1 c_{H_2} \sum_{i=0}^{n-1} \mathbf{E} \left(K(n^\alpha B_i^{H_1}) \right) (-|i+1-t|^{2H_2} + |i-t|^{2H_2} + |i+1|^{2H_2} - i^{2H_2}) \end{aligned}$$

where c_1 is the constant upper bound of the derivative of G_t and c_{H_2} is a constant depending only on H_2 . Since for every fixed $t > 0$ the function $(-|i+1-t|^{2H_2} + |i-t|^{2H_2} + |i+1|^{2H_2} - i^{2H_2}) =$ behaves, modulo a constant, as i^{2H_2-2} and since the order of the expectation of $K(n^\alpha B_i^{H_1})$ is the same as that of $n^{-\alpha} i^{-H_1}$ it is clear that T_1 converges to zero as $n \rightarrow \infty$.

Finally, we will show that T_3^* converges to zero. Note that the term T_3^* can be expressed as follows

$$T_3^* = \lambda_n^2 \sum_{i=0}^{n-1} \mathbf{E} \left(K^2(n^\alpha B_i^{H_1}) \int_i^{i+1} e^{i_0 \lambda_n \eta_s^{(\lambda_n)}} \left(\frac{1}{2} - H_2(s-i)^{2H_2-1} \right) ds \cdot e(\lambda, n) g_N \right).$$

At this point, we will again apply the Itô formula for $e^{i\lambda_n \eta_s^{(\lambda_n)}}$. It implies that

$$\begin{aligned}
T_3^* &= \mathbf{E} \left(\lambda_n^2 \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) \int_i^{i+1} \left(\frac{1}{2} - H_2(s-i)^{2H_2-1} \right) ds \cdot e(\lambda, n) g_N \right) \\
&\quad + \mathbf{E} \left(i_0 \lambda_n^3 \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) \int_i^{i+1} \left(\frac{1}{2} - H_2(s-i)^{2H_2-1} \right) \right. \\
&\quad \left. \int_0^s e^{i_0 \lambda_n \eta_u^{(\lambda_n)}} K(n^\alpha B_{[u]}^{H_1}) dB_u^{H_2} ds \cdot e(\lambda, n) g_N \right) \\
&\quad + \mathbf{E} \left(\frac{1}{2} \lambda_n^4 \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) \int_i^{i+1} \left(\frac{1}{2} - H_2(s-i)^{2H_2-1} \right) \right. \\
&\quad \left. \int_0^s e^{i_0 \lambda_n \eta_u^{(\lambda_n)}} K^2(n^\alpha B_{[u]}^{H_1}) du ds \cdot e(\lambda, n) g_N \right) \\
&\quad - \mathbf{E} \left(\lambda_n^4 \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) \int_i^{i+1} \left(\frac{1}{2} - H_2(s-i)^{2H_2-1} \right) \right. \\
&\quad \left. \times H_2(2H_2-1) \int_0^s e^{i_0 \lambda_n \eta_u^{(\lambda_n)}} K(n^\alpha B_{[u]}^{H_1}) \int_0^u K(n^\alpha B_{[v]}^{H_1}) |u-v|^{2H_2-2} dv du ds \cdot e(\lambda, n) g_N du \right) \\
&= b^{(1)} + b^{(2)} + b^{(3)} + b^{(4)}. \tag{39}
\end{aligned}$$

The first summand $b^{(1)}$ vanishes because the integral

$$\int_i^{i+1} \left(\frac{1}{2} - H_2(s-i)^{2H_2-1} \right) ds$$

vanishes. The second summand $b^{(2)}$ goes to zero as $n \rightarrow \infty$ using exactly the same argument as for the convergence of T_1^* . Concerning the third summand, $b^{(3)}$, using (38) and the fact that $|g_N| = 1$, we get

$$\begin{aligned}
b^{(3)} &\leq \mathbf{E} \left(\lambda_n^4 \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) \int_i^{i+1} \underbrace{\left| \frac{1}{2} - H_2(s-i)^{2H_2-1} \right|}_{\leq 1} \int_0^s K^2(n^\alpha B_{[u]}^{H_1}) du ds \right) \\
&\leq \mathbf{E} \left(\frac{1}{2} \lambda_n^4 \sum_{i=0}^{n-1} K^2(n^\alpha B_i^{H_1}) \int_i^{i+1} \left(\sum_{j=0}^{i-1} K^2(n^\alpha B_j^{H_1}) + K^2(n^\alpha B_i^{H_1}) \underbrace{(s-i)}_{\leq 1} \right) ds \right) \\
&\leq \mathbf{E} \left(\lambda_n^4 \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K^2(n^\alpha B_i^{H_1}) K^2(n^\alpha B_j^{H_1}) \right) + \mathbf{E} \left(\lambda_n^4 \sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K^4(n^\alpha B_i^{H_1}) \right).
\end{aligned}$$

The second term goes to zero because $\mathbf{E} \left(\sum_{i=0}^{n-1} \sum_{j=0}^{i-1} K^4 \left(n^\alpha B_i^{H_1} \right) \right)$ behaves as $n^{-\alpha-H_1+1}$ and the first term goes to zero because the non-diagonal term is dominated by the diagonal term. Analogously to the convergence of T_2^* , the last summand in (39) converges to zero. This completes the proof. \blacksquare

References

- [1] S. Berman (1973): *Local nondeterminism and local times of Gaussian processes*. Indiana Univ. Math. J. 23, 6994.
- [2] F. Biagini, Y. Hu, B. Oksendal and T. Zhang (2008): *Stochastic calculus for fractional Brownian motion and applications*. Springer.
- [3] L. Coutin, D. Nualart and C.A. Tudor (2001): *The Tanaka formula for the fractional Brownian motion*. Stochastic Proc. Applic., 94(2), 301-315.
- [4] M. Eddahbi, R. Lacayo, J.L. Sole, C.A. Tudor and J. Vives (2001): *Regularity of the local time for the d-dimensional fractional Brownian motion with N-parameters*. Stochastic Analysis and Applications, 23(2), 383-400.
- [5] D. Geman and J. Horowitz (1980): *Occupation densities*. Ann. Probab. 8, 167.
- [6] Yaozhong Hu and B. Oksendal (2002): *Chaos expansion of local times of fractional Brownian motions*. Stochastic Analysis and Applications, 20 (4), 815-837
- [7] P. Imkeller and P. Weisz (1994): *The asymptotic behavior of local times and occupation integrals of the N -parameter Wiener process in \mathbb{R}^d* . Prob. Th. Rel. Fields, 98 (1), 47-75.
- [8] H.A. Karlsen and D. Tjøstheim (2001): *Nonparametric estimation in null recurrent time series*. The Annals of Statistics, 29, 372-416.
- [9] H.A. Karlsen, T. Mykklebust and D. Tjøstheim (2007): *Non parametric estimation in a nonlinear cointegrated model*. The Annals of Statistics, 35, 252-299.
- [10] D. Nualart (2006): *The Malliavin Calculus and Related Topics*. Second edition, Springer.
- [11] D. Nualart and J. Vives (1992): *Smoothness of Brownian local times and related functionals*. Potential Analysis, 1(3), 257-263.
- [12] D. Nualart and J. Vives (2002): *Chaos expansiun and local times*. Publicacions Matematiques, 36, 827-836.
- [13] J.Y. Park and P.C.B. Phillips (2001): *Nonlinear regression with integrated time series*. Econometrica, 74, 117-161.

- [14] P.C.B. Phillips (1988): *Regression theory for near-integrated time series*. Econometrica, 56, 1021-1044.
- [15] Q. Wang and P. Phillips (2009): *Asymptotic Theory for the local time density estimation and nonparametric cointegrated regression*. Econometric Theory, 25, 710-738.
- [16] Q. Wang and P. Phillips (2009): *Structural Nonparametric cointegrating regression*. Econometrica, 77(6), 1901-1948.
- [17] L. Yan, J. Liu and X. Yang (2009): *Integration with respect to fractional local time with Hurst index $\frac{1}{2} < H < 1$* . Potential Analysis, 30, 115-138.